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NEW CALIBRATION OF FINITE POPULATION MEAN OF COMBINED RATIO ESTIMATORS IN STRATIFIED RANDOM SAMPLING

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ABSTRACT

This study deals with using calibration estimation approaches to modified the combined ratio estimator in stratified random sampling. Calibration distance measures with their associate constraints were used to modify combined ratio estimator. In stratified random sampling, new sets of optimum calibration weights are created and used to obtain new calibration estimators of population mean. Empirical study through simulation was conducted to look into the efficiency of the suggested estimators obtained. The suggested calibration estimators are more efficient than other existing estimators investigated in the study, according to the findings.

Keywords: Weights for calibration, Combined Ratio, Estimators, Mean Squared Error, Stratified Sampling

INTRODUCTION

Estimation of calibration by incorporating the known population characteristics of auxiliary variables, the original design weights can be adjusted. In stratified random sampling, a calibration approach is used to determine the optimum stratum weights for enhancing the precision of survey estimates of population parameters. Deville and Sarndal established the approach of estimate by calibration in survey sampling in 1992. The ideal is to use auxiliary information to improve parameter estimates for the population of interest. Following Deville and Sarndal (1992), many researchers have investigated survey sample design calibration estimation utilizing various calibration constraints on auxiliary variables. Singh et al. (1998) was the first researcher that extended calibration approach to stratified sampling design. In stratified sampling, Tracy et al. (2003), Singh (2003), Kim *et al.* (2007), Clement and Enang (2015) and Ozgul (2018) used calibration estimation in ratio-type estimators. Rao et al. (2012) used different distance measures

with two auxiliary variables in stratified sampling to construct a multivariate calibration estimate for the population mean. Based on distinct calibration constraints of auxiliary information, Koyuncu and Kadilar (2016) proposed calibration estimators for estimating the population mean in stratified sampling. The existing calibration estimators such as Singh et al. (1998), Tracy et al. (2003), Kim et al. (2007), Rao et al. (2012), Koyuncu and Kadilar (2013), Clement and Enang (2015), Koyuncu and Kadilar (2016), Clement (2017) and Ozgul (2018) adjusted the stratum weight of unbiased estimators utilizing auxiliary variables in stratified sampling to increase efficiency, favoring and more accurately representing strata with large sizes. To overcome this challenge, the strata sample means of the study variable are modified using ratio function of population and sample means of auxiliary variable within the strata using calibration techniques to produce more efficient estimate. Muili, et al. (2019).

Notations and Review of Existing Estimators

Take a look at a finite population ψ N elements, $\psi = \{\psi_1, \psi_2, \psi_3, ..., \psi_N\}$ consists of L strata with N_h units in the h th stratum from which a simple random of size n_h can be generated from the population using SRSWOR. Total Population

size
$$N = \sum_{h=1}^{L} N_h$$
, sample size $n = \sum_{h=1}^{L} n_h$, y_{hi} , $i = 1, 2, ..., N_{hi}$ and x_{hi} , $i = 1, 2, ..., N_{hi}$ of study variable y and y_{hi} , $y_{$

X auxiliary variable. Let $W_h = N_h/N$ the weights of the strata, $\overline{Y}_h = N^{-1} \sum_{i=1}^{n_h} y_{hi}$ and $\overline{y}_h = n^{-1} \sum_{i=1}^{n_h} y_{hi}$ are the population and sample means respectively for the study variables.

The classic estimator of population mean in stratified sampling, according to Cochran (1977), is: $\overline{y}_{st} = \sum_{h=1}^{L} W_h \overline{y}_h$ (2.0)

$$V\left(\overline{y}_{st}\right) = \sum_{h=1}^{L} W_h^2 \left(\frac{1 - f_h}{n_h}\right) s_{hy}^2 \tag{2.1}$$

where
$$s_{hy}^2 = (n_h - 1)^{-1} \sum_{h=1}^{n_h} (y_{hi} - \overline{y}_h)^2$$
, $f_h = n_h / N_h$

A combined ratio estimator was proposed by Hansen et al. (1946).

$$\overline{y}_{st}^{RC} = \frac{\overline{y}_{st}}{\overline{x}_{st}} \overline{X}$$
 (2.2)

where
$$\overline{x}_{st} = \sum_{i=1}^{n_h} W_h \overline{x}_h$$
 and $\overline{y}_{st} = \sum_{i=1}^{n_h} W_h \overline{y}_h$

The combined ratio estimator's variance is

$$V(\bar{y}_{st}^{RC}) = \sum_{h=1}^{L} W_h^2 \gamma_h \left(S_{yh}^2 + R^2 S_{xh}^2 - 2R S_{yxh}^2 \right)$$
 (2.3)

where
$$R = \frac{\overline{Y}}{\overline{X}}$$

Lata *et al.* (2017) used a distance function of the sum of weighted squared deviation of calibrated and strata weights to build a calibration estimator for estimating population mean under stratified random sampling:

$$\overline{y}_{st}^{LA} = \sum_{h=1}^{L} \Omega_h^{LA} \overline{y}_h \tag{2.4}$$

$$Z = \sum_{h=1}^{L} S_{hx}^{2} \left(Q_{h} \right)^{-1} \left(\Omega_{h}^{LA} - W_{h} \right)^{2}$$
 (2.5)

where Ω_h^{LA} are the calibrated weights and S_{hx}^2 is the auxiliary variable's variance in the hth stratum, subject to the calibration constraint.

$$\sum_{h=1}^{L} \Omega_h^{LA} \overline{x}_{_{1h}} = \overline{X} \tag{2.6}$$

The calibrated weights and the estimator of (2.4) are obtained as show in (2.7) and (2.8) respectively.

$$\Omega_{h}^{LA} = W_{h} + \frac{\overline{x}_{h} Q_{h} \left(S_{hx}^{2}\right)^{-1}}{\sum_{h=1}^{L} \overline{x}_{h}^{2} Q_{h} \left(S_{hx}^{2}\right)^{-1}} \left(\overline{X} - \sum_{h=1}^{L} W_{h} \overline{x}_{h}\right)$$
(2.7)

$$\overline{y}_{st}^{LA} = \sum_{h=1}^{L} W_h \overline{y}_h + \hat{\beta} \left(\overline{X} - \sum_{h=1}^{L} W_h \overline{x}_h \right)$$
(2.8)

where
$$\hat{\beta} = \frac{\sum_{h=1}^{L} \overline{x}_{h} \overline{y}_{h} Q_{h} \left(S_{hx}^{2}\right)^{-1}}{\sum_{h=1}^{L} \overline{x}_{h}^{2} Q_{h} \left(S_{hx}^{2}\right)^{-1}}$$

The calibrated estimator's variance is given by

$$V(\bar{y}_{st}^{LA}) = \sum_{h=1}^{L} \Omega_h^{LA^2} \frac{(1-f_h)}{n_h} s_{eh}^2$$
(2.9)

where $s_{eh}^2 = \left(n_h - 1\right)^{-1} \sum_{i=1}^{nh} e_{hi}^2$ is the mean square of the hth stratum sample and $e_{hi} = \left(y_{hi} - \overline{y}_h\right) - \hat{\beta}\left(x_{hi} - \overline{x}_h\right)$.

MATERIALS AND METHODS

Suggested Estimator

Having studied Lata *et al.* (2017) calibration estimator, and motivated by his work. A new set of calibration estimators are suggested as:

$$\overline{y}_{st}^{MJ} = \sum_{h=1}^{L} \Omega_h^{MJ} \overline{y}_h \tag{3.0}$$

$$Z^* = \sum_{h=1}^{L} S_{hx}^2 \left(Q_h \right)^{-1} \left(\Omega_h^{MJ} - W_h^* \right)^2$$
 (3.1)

where $W_h^* = W_h \overline{X} / \sum_{h=1}^L W_h \overline{X}_h$, Ω_h^{MJ} are the new calibrated weights and S_{hx}^2 is the variance of auxiliary variable in hth

stratum. The form of the suggested estimators is determined by Q_{h} , which is a selected weight.

Subject to the calibration constraint

$$\sum_{h=1}^{L} \Omega_h^{MJ} \overline{x}_h = \sum_{h=1}^{L} W_h^* \overline{X}_h$$
 (3.2)

The Lagrange Multiplier function L of \overline{y}_{st}^{MJ} is defined as follows to get new calibrated weights $\left(\Omega_h^{MJ}\right)$ for the suggested estimator $\left(\overline{y}_{st}^{MJ}\right)$:

$$L = \sum_{h=1}^{L} S_{xh}^{2} \left(Q_{h} \right)^{-1} \left(\Omega_{h}^{MJ} - W_{h}^{*} \right)^{2} - 2\lambda \left(\sum_{h=1}^{L} \Omega_{h}^{MJ} \overline{x}_{h} - \sum_{h=1}^{L} W_{h}^{*} \overline{X}_{h} \right)$$
(3.3)

where λ is Lagrange's multiplier, differentiate (3.3) in relation to Ω_h^{MJ} , λ and are equal to zero to obtained (3.4), and (3.5) as follows:

$$\Omega_h^{MJ} = W_h^* + \lambda Q_h \overline{x}_h \left(S_{xh}^2 \right)^{-1} \tag{3.4}$$

$$\sum_{h=1}^{L} \Omega_h^{MJ} \bar{x}_h - \sum_{h=1}^{L} W_h^* \bar{X}_h = 0$$
 (3.5)

Substitute (3.4) in (3.5), the results are obtained as:

$$\lambda \sum_{h=1}^{L} Q_h \overline{x}_h^2 \left(S_{xh}^2 \right)^{-1} = \sum_{h=1}^{L} W_h^* \overline{X}_h - \sum_{h=1}^{L} W_h^* \overline{x}_h \tag{3.6}$$

Make λ the subject of formula, obtained as

$$\lambda = \frac{\sum_{h=1}^{L} W_h^* \overline{X}_h - \sum_{h=1}^{L} W_h^* \overline{X}_h}{\sum_{h=1}^{L} Q_h \overline{X}_h^2 \left(S_{xh}^2\right)^{-1}}$$
(3.7)

When (3.7) is substituted for (3.4), the calibrated weights become:

$$\Omega_{h}^{MJ} = W_{h}^{*} + \left(Q_{h}\overline{x}_{h}\left(S_{xh}^{2}\right)^{-1}\right) \left(\frac{\sum_{h=1}^{L} W_{h}^{*}\overline{X}_{h} - \sum_{h=1}^{L} W_{h}^{*}\overline{x}_{h}}{\sum_{h=1}^{L} Q_{h}\overline{x}_{h}^{2}\left(S_{xh}^{2}\right)^{-1}}\right)$$
(3.8)

Substituting (3.8) in (3.0), obtain the new combined calibration estimator $\left(\overline{y}_{st}^{MJ}\right)$ as:

$$\bar{y}_{st}^{MJ} = \sum_{h=1}^{L} W_h^* \bar{y} + \hat{\beta}_1 \sum_{h=1}^{L} W_h^* \left(\bar{X}_h - \bar{x}_h \right)$$
(3.9)

Substituting $W_h^* = W_h \overline{X} / \sum_{h=1}^L W_h \overline{x}_h$ in (3.9), gives

$$\overline{y}_{st}^{MJ} = \overline{X} \left(\sum_{h=1}^{L} W_h \overline{y} \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} + \hat{\beta}_1 \overline{X} \left(\sum_{h=1}^{L} W_h \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} \left(\overline{X}_h - \overline{x}_h \right)$$
(3.11)

where
$$\hat{\beta}_{1} = \frac{\sum_{h=1}^{L} Q_{h} \overline{x}_{h} \overline{y} \left(S_{xh}^{2}\right)^{-1}}{\sum_{h=1}^{L} Q_{h} \overline{x}_{h}^{2} \left(S_{xh}^{2}\right)^{-1}}$$

Setting $Q_h = 1$, $Q_h = \overline{x}_h^{-1}$, and $Q_h = \left(S_{xh}^2\right)^{-1}$, we have the following new set of calibration combined ratio estimations respectively:

$$\overline{y}_{st11}^{MJ} = \overline{X} \left(\sum_{h=1}^{L} W_h \overline{y} \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} + \hat{\beta}_{11} \overline{X} \left(\sum_{h=1}^{L} W_h \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} \left(\overline{X}_h - \overline{x}_h \right)$$
(3.12)

$$\overline{y}_{st12}^{MJ} = \overline{X} \left(\sum_{h=1}^{L} W_h \overline{y} \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} + \hat{\beta}_{12} \overline{X} \left(\sum_{h=1}^{L} W_h \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} \left(\overline{X}_h - \overline{x}_h \right)$$
(3.13)

$$\overline{y}_{st13}^{MJ} = \overline{X} \left(\sum_{h=1}^{L} W_h \overline{y} \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} + \hat{\beta}_{13} \overline{X} \left(\sum_{h=1}^{L} W_h \right) \left(\sum_{h=1}^{L} W_h \overline{x}_h \right)^{-1} \left(\overline{X}_h - \overline{x}_h \right)$$
(3.14)

where
$$\hat{\beta}_{11} = \frac{\sum_{h=1}^{L} \overline{x}_h \overline{y} \left(S_{xh}^2\right)^{-1}}{\sum_{h=1}^{L} \overline{x}_h^2 \left(S_{xh}^2\right)^{-1}}$$
, $\hat{\beta}_{12} = \frac{\sum_{h=1}^{L} \overline{y} \left(S_{xh}^2\right)^{-1}}{\sum_{h=1}^{L} \overline{x}_h \left(S_{xh}^2\right)^{-1}}$ and $\hat{\beta}_{13} = \frac{\sum_{h=1}^{L} \overline{x}_h \overline{y} \left(S_{xh}^4\right)^{-1}}{\sum_{h=1}^{L} \overline{x}_h^2 \left(S_{xh}^4\right)^{-1}}$

Table 1: Populations Involved in the Empirical Research

Population	Auxiliary variable $^{\mathcal{X}}$	Study variable ^y
I	$x_h \approx \exp(\theta_h), \theta_1 = 5, \theta_2 = 6,$	
	$\theta_3 = 4, h = 1, 2, 3$	
II	$x_h \approx gamma(\theta_h, \eta_h), \theta_1 = 3, \eta_1 = 2,$	$y_{hi} = \alpha_h x_{hi}^j + \xi_{hi}, \ \alpha_{1h} = E(x_h),$
	$\theta_2 = 3, \eta_2 = 1, \theta_3 = 3, \eta_3 = 3,$	$\alpha = 1.0, \ \xi_h \approx N(0,1), h = 1,2,3. \ j = 2,3,4$
III	$x_h \approx \operatorname{chisq}(\theta_h), \theta_1 = 5, \theta_2 = 6,$	
	$\theta_3 = 4, h = 1, 2, 3$	

Table 2: MSE and PRE of Some Existing and Suggested Estimators Using Population I

Estimator	$y_{hi} = \alpha_h x_{hi}^2 + \xi_{hi}$		
	MSE	PRE	
\overline{y}_{st}	4.400796	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	1.685452	261.1048	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st11}^{LA}$	1.93655	227.2493	
Lata et al. (2017) \overline{y}_{st12}^{LA}	2.028274	216.9725	
Lata et al. (2017) \overline{y}_{st13}^{LA}	2.149103	204.7736	
Suggested Estimator \overline{y}_{st11}^{MJ}	0.8508302	517.2355	
Suggested Estimator \overline{y}_{st12}^{MJ} Suggested Estimator	0.8646045	508.9953	
Suggested Estimator \overline{y}_{st13}^{MJ} Suggested Estimator	0.8871836	496.0412	
Suggested Estimator 3 st13			

Table 3: MSE and PRE of Some Existing and Suggested Estimators Using Population I

Estimator	$y_{hi} = lpha_h x_{hi}^3 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	556.1768	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	354.32	156.9702	
Lata et al. (2017) \overline{y}_{st11}^{LA}	398.3078	139.6349	
Lata et al. (2017) \overline{y}_{st12}^{LA} Lata et al. (2017)	412.5612	134.8107	
Lata et al. (2017) \overline{y}_{st13}^{LA}	431.4772	128.9006	
Suggested Estimator \overline{y}_{st11}^{MJ}	271.1696	205.1029	
Suggested Estimator \overline{y}_{st12}^{MJ} Suggested Estimator \overline{y}_{st12}^{MJ}	277.8351	200.1823	
Suggested Estimator \overline{y}_{st13}^{MJ} Suggested Estimator \overline{y}_{st13}^{MJ}	287.1083	193.7167	
Suggested Estimator 5 st13			

Table 4: MSE and PRE of Some Existing and Suggested Estimators Using Population I

Estimator	$y_{hi} = \alpha_h x_{hi}^4 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	94055.7	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	71867.57	130.8736	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st11}^{LA}$	79068.55	118.9546	
Lata et al. (2017) \overline{y}_{st12}^{LA}	81246.93	115.7652	
Lata et al. (2017) \overline{y}_{st13}^{LA}	84021.24	111.9428	
Suggested Estimator \overline{y}_{st11}^{MJ}	63159.03	148.9188	
Suggested Estimator \overline{y}_{st12}^{MJ}	64357.91	146.1447	
	65918.68	142.6844	
Suggested Estimator \mathcal{Y}_{st13}^{-1}			

Table 5: MSE and PRE of Some Existing and Suggested Estimators Using Population II

Estimator	$y_{hi} = \alpha_h x_{hi}^2 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	0.7243471	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	0.3669404	197.4018	
Lata et al. (2017) \overline{y}_{st11}^{LA}	0.4105255	176.4439	
Lata et al. (2017) \overline{y}_{s112}^{LA} Lata et al. (2017)	0.4556056	158.9855	
Lata et al. (2017) \overline{y}_{st13}^{LA} Lata et al. (2017)	0.4856501	149.15	
Suggested Estimator $\overline{\mathcal{Y}}_{st11}^{MJ}$	0.2231978	324.5314	
Suggested Estimator \overline{y}_{st12}^{MJ} Suggested Estimator \overline{y}_{st12}^{MJ}	0.239531	302.4022	
Suggested Estimator \overline{y}_{st13}^{MJ} Suggested Estimator	0.2519612	287.4836	
Suggested Estimator St13			

Table 6: MSE and PRE of Some Existing and Suggested Estimators Using Population II

Estimator	$y_{hi} = \alpha_h x_{hi}^3 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	37.77247	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	26.38607	143.1531	
Lata et al. (2017) \overline{y}_{st11}^{LA}	29.33956	128.7424	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st12}^{LA}$	31.85461	118.5777	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st13}^{LA}$	33.45616	112.9014	
Suggested Estimator \overline{y}_{st11}^{MJ}	20.8789	180.9122	
Suggested Estimator $\overline{\mathcal{Y}}_{st12}^{MJ}$	22.42852	168.4127	
-MJ	23.44996	161.0769	
Suggested Estimator y_{st13}			

Table 7: MSE and PRE of Some Existing and Suggested Estimators Using Population II

Estimator	$y_{hi} = \alpha_h x_{hi}^4 + \xi_{hi}$	
	MSE	PRE
\overline{y}_{st}	2601.545	100
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	2093.923	124.2426
Lata et al. (2017) $\overline{\mathcal{Y}}_{st11}^{LA}$	2274.021	114.4028
Lata et al. (2017) $\overline{\mathcal{Y}}_{st12}^{LA}$	2408	108.0376
Lata et al. (2017) \overline{y}_{st13}^{LA}	2490.562	104.4562
Suggested Estimator \overline{y}_{st11}^{MJ}	1863.71	139.5896
Suggested Estimator $\frac{\overline{y}_{st12}^{MJ}}{\overline{y}_{st12}^{SI}}$	1955.269	133.053
-M1	2012.79	129.2507
Suggested Estimator y_{st13}^{m}		

Table 8: MSE and PRE of Some Existing and Suggested Estimators Using Population III

Estimator	$y_{hi} = \alpha_h x_{hi}^2 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	3.505272	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	1.267897	276.4634	
Combined ratio \overline{y}_{st11}^{LA} Lata et al. (2017)	1.404499	249.5745	
Lata et al. (2017) \overline{y}_{st12}^{LA} Lata et al. (2017)	1.596074	219.6185	
Lata et al. (2017) \overline{y}_{st13}^{LA} Lata et al. (2017)	1.674268	209.3615	
Suggested Estimator $\overline{\mathcal{Y}}_{st11}^{MJ}$	0.636964	550.3093	
Suggested Estimator $\frac{\nabla MJ}{\nabla MJ}$	0.6503848	538.9537	
Suggested Estimator \overline{y}_{st12}^{MJ} Suggested Estimator \overline{y}_{st13}^{MJ}	0.6660098	526.3094	
Suggested Estimator \mathcal{Y}_{st13}			

Table 9: MSE and PRE of Some Existing and Suggested Estimators Using Population III

Estimator	$y_{hi} = \alpha_h x_{hi}^3 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	342.7701	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	210.0028	163.2217	
Lata et al. (2017) \overline{y}_{st11}^{LA}	225.9159	151.7246	
Lata et al. (2017) \overline{y}_{st12}^{LA} Lata et al. (2017)	245.4806	139.6323	
Lata et al. (2017) \overline{y}_{st13}^{LA} Lata et al. (2017)	253.9479	134.9765	
Suggested Estimator \overline{y}_{st11}^{MJ}	151.2407	226.6388	
Suggested Estimator $\frac{\nabla MJ}{\nabla MJ}$	159.5417	214.8467	
Suggested Estimator \overline{y}_{st12}^{MJ} Suggested Estimator \overline{y}_{st13}^{MJ}	163.5477	209.5842	
Suggested Estimator y_{st13}			

Table 10: MSE and PRE of Some Existing and Suggested Estimators Using Population III

Estimator	$y_{hi} = \alpha_h x_{hi}^4 + \xi_{hi}$		
	MSE	PRE	
$\overline{\mathcal{Y}}_{st}$	44896.74	100	
Combined ratio $\overline{\mathcal{Y}}_{st}^{RC}$	34022.95	131.9602	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st11}^{LA}$	35868.74	125.1695	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st12}^{LA}$	37950.3	118.304	
Lata et al. (2017) $\overline{\mathcal{Y}}_{st13}^{LA}$	38854.48	115.551	
Suggested Estimator $\overline{\mathcal{Y}}_{st11}^{MJ}$	28877.71	155.4719	
Suggested Estimator \overline{y}_{st12}^{MJ}	29974.4	149.7836	
Suggested Estimator \overline{y}_{st13}^{MJ}	30470.31	147.3459	
Suggested Estimator * 3715			

RESULTS AND DISCUSSION

Tables 2-10 show the Mean Square Error (MSE) and Percentage Relative Efficiency (PRE) of the suggested combined calibration estimators and other estimators using simulated data. The results revealed that the suggested combined calibration estimators have minimum MSE compared to the traditional estimators and other estimators considered under stratified random sampling.

CONCLUSION

The new combined stratified calibration estimators for estimating population mean is suggested with a set of constraint which minimized a given chi-squared distance measure. New calibration weights are developed. The simulation analysis reveals that the suggested calibration estimators outperform other estimators in the study (with lower MSE and higher PRE). As a result, we conclude that the suggested calibration estimators are more efficient than the traditional estimator, combined ratio estimator, and Lata *et al.* (2017) estimators in estimating the population mean of the study variable.

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